# Nonparametric approaches to assessing variable importance using health data

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Variable importance is the quantification of the contribution of a feature (or group of features) toward a learning task.

- Predicting an outcome
- Classifying an observation
- Determining a treatment rule



## Example 1: suicide prevention

**Motivation:** Clinical risk prediction models for suicide and self-harm use data from EHR and insurance claims available at the time of a healthcare visit.

 Data reflecting recent events (prescriptions, diagnoses, self-harm events) may not be available in real time.



Weckstein et al. "Data lag in a large open and closed claims dataset: Navigating the completeness-timeliness tradeoff." Wolock et al. "Importance of variables from different time frames for predicting self-harm using health system data." **Motivation:** Identify binding and neutralizing antibodies that predict protection against SARS-CoV-2 infection.

▶ Hypothesis generation for a formal correlates of immunity analysis



Gilbert et al. "A Covid-19 Milestone Attained — A Correlate of Protection for Vaccines." Hoffman et al. "Correlates of Protection Against SARS-CoV-2 Infection in Children." From 2016 to 2019, the HIV Vaccine Trials Network conducted a trial to investigate the efficacy of a recombinant canarypox vaccine targeted at HIV-1 subtype C (prevalent in sub-Saharan Africa) in adults aged 18-35.



**Secondary objective:** learn about factors that predict risk of HIV acquisition.

Risk models include:

Age	Prevalent STIs
Sex	Behavioral characteristics
BMI	Partner characteristics
Housing	Geography

#### What is the relative importance of these features?

▶ In particular, how much do we gain from relatively "expensive" predictors?

Gray et al. "Vaccine efficacy of ALVAC-HIV and bivalent subtype C gp120-MF59 in adults." Wolock, Gilbert, Simon, Carone. "Assessing variable importance in survival analysis using machine learning."

#### Types of variable importance



#### Types of variable importance



#### Predictiveness

The ideal data unit is  $(X, T) \sim \mathbb{P}_0$ , which lies in a nonparametric model  $\mathcal{M}$ .

- ► *T* is the outcome of interest.
- $X = (X_1, \ldots, X_p)$  are the features, predictors, or covariates.

We use f to denote a generic prediction function.

▶ The performance of f under sampling from distribution  $\mathbb{P}$  is quantified by the predictiveness  $\mathbb{V}(f, \mathbb{P})$ , e.g., minus mean squared error.

$$\begin{pmatrix} X_1 \\ \vdots \\ X_p \end{pmatrix} \xrightarrow{f} f(X) \xrightarrow{\text{compare to } T} -\{f(X) - T\}^2 \xrightarrow{\text{summarize over } \mathbb{P}} -E_{\mathbb{P}}\left[\{f(X) - T\}^2\right]$$

In this case, 
$$\mathbb{V}(f,\mathbb{P}) = -E_{\mathbb{P}}\left[\{f(X) - T\}^2\right]$$
.

Note: Throughout this talk, 'blackboard' font represents the ideal data world, regular font the observed data world.

$$\mathbb{P} \longleftrightarrow \mathsf{P}_{\mathsf{obs. data}}$$

#### Focusing on survival analysis

In each of the settings we consider, the outcome of interest T is the time between an initiating event and a terminating event.



For now, we work on the scale of follow-up time, so the initiating event occurs at time t = 0 for all individuals.

Some common predictiveness measures in the survival setting:

- ► <u>AUC at time  $\tau$ </u>:  $\mathbb{V}(f, \mathbb{P}_0) = \mathbb{P}_0 \{ f(X_1) > f(X_2) \mid T_1 \leq \tau, T_2 > \tau \}$ (Heagerty and Zheng, 2005)
- Brier score at time  $\tau$ :  $\mathbb{V}(f, \mathbb{P}_0) = -E_{\mathbb{P}_0}[\{f(X) \mathbb{1}(T \leq \tau)\}^2]$ (Brier, 1950)
- <u>C-index</u>:  $\mathbb{V}(f, \mathbb{P}_0) = \mathbb{P}_0 \{ f(X_1) > f(X_2) \mid T_1 \leq T_2 \}$ (Harrell, 1982)

## The key role of (informative) censoring



Cumulative Incidence of LTFU\* over Study Time by Treatment Enrolled Participants

Two types of censoring:

- 1 Study termination (administrative)
- 2 Participant drop-out



#### The current landscape

What methods exist for estimating variable importance under right censoring?

- **Algorithmic:** Parametric/semiparametric models, e.g., coefficients from a Cox model (Cox, 1972).
  - Dependent on correct model specification
  - Difficult to compare across features
  - Not immediately clear how to handle interactions, groups of features, correlated features, etc.
  - Interpretation not necessarily linked to predictiveness

2 Algorithm-agnostic: Permutation (Breiman, 2001; Fisher et al., 2019).



#### Perils of permutation importance

- ► Lacking methods for estimation and inference under informative censoring.
- "Extrapolation bias": For correlated features, this approach requires generating predictions in regions with little (or even zero) probability mass under the joint distribution of the features (Hooker et al., 2021; Wang et al., 2024; Verdinelli and Wasserman, 2024).



**Exclusion importance** — comparing models trained using different subsets of features — does not suffer this same issue.

## The oracle prediction function

**Goal:** assess the importance of  $X_s$ , where  $s \subset \{1, ..., p\}$ , relative to the full predictor vector X. In the exclusion paradigm, we consider two special prediction functions:

 $f_0 := \operatorname{argmax}_{f \in \mathcal{F}} \mathbb{V}(f, \mathbb{P}_0)$ 

 The 'full' oracle, optimal in an unrestricted class of prediction functions

$$f_{0,s} := \operatorname{argmax}_{f \in \mathcal{F}_s} \mathbb{V}(f, \mathbb{P}_0)$$

 The 'reduced' oracle, optimal in the class of prediction functions that does not use X<sub>s</sub>



We define the importance of  $X_s$  relative to X as  $\mathbb{V}(\mathbb{f}_0, \mathbb{P}_0) - \mathbb{V}(\mathbb{f}_{0,s}, \mathbb{P}_0)$ .

- ▶ Rather than this subtractive notion, we could also consider an additive approach where X<sub>s</sub> is added to a 'base' set of predictors.
- ► It may also be of interest to normalize by V(f<sub>0</sub>, P<sub>0</sub>); estimation and inference can be handled using the delta method.

The parameter  $\mathbb{V}(\underline{f}_0, \mathbb{P}_0) - \mathbb{V}(\underline{f}_{0,s}, \mathbb{P}_0)$  is a sensible quantification of the importance of  $X_s$  relative to X.

Taking this as our parameter of interest, we next focus on

- identification;
- 2 estimation;
- 3 inference.

Due to right censoring, the predictiveness measure  $\ensuremath{\mathbb{V}}$  is not a functional of the observed data distribution.

Many predictiveness measures have a common form that we can exploit:

$$\mathbb{V}(f_0, \mathbb{P}_0) = E_{\mathbb{P}_0} \left( \omega \left[ \{ f_0(X_1), T_1 \}, \dots, \{ f_0(X_m), T_m \} \right] \right) \\ = \int \cdots \int \omega \left[ \{ f_0(x_1), t_1 \}, \dots, \{ f_0(x_m), t_m \} \right] \prod_{j=1}^m \mathbb{H}_0(dx_j, dt_j).$$

where  $\omega$  is a known kernel function and  $\mathbb{H}_0$  is the joint cdf of (X, T) under  $\mathbb{P}_0$ .

With a slight abuse of notation, we write  $\mathbb{V}(\underline{f}_0, \mathbb{P}_0)$  as  $\mathbb{V}(\underline{f}_0, \mathbb{H}_0)$ .

If T and C are independent within strata defined by X, then  $\mathbb{H}_0$  is identified by the observed data distribution.

In contrast, many existing methods for evaluating predictiveness with censored data make a stronger marginal independence assumption, simplifying estimation.

Under conditionally independent censoring, the joint cdf  $\mathbb{H}_0$  can be identified pointwise — for some values of  $(x_0, t_0)$  — as

$$\mathbb{H}_0(x_0, t_0) = \int \mathbb{1}(x \leq x_0) G_0(t_0 \mid x) Q_0(dx) := H_0(x_0, t_0) ,$$

with

- ► G<sub>0</sub> the product-integral mapping applied to the conditional cumulative hazard function of T given X under P<sub>0</sub>;
- $Q_0$  the marginal cdf of X under  $\mathbb{P}_0$ .

The identified parameter  $V(f_0, H_0) - V(f_{0,s}, H_0)$  depends on the nuisance functions  $(f_0, f_{0,s}, G_0)$ .

► We want to use flexible (machine learning) methods, e.g., random survival forests (Ishwaran et al., 2008), survival Super Learner (Westling et al., 2024), survival stacking (Wolock et el., 2024).

A first attempt:

- **E**stimate the nuisance functions  $(f_0, f_{0,s}, G_0)$  using machine learning estimators  $(f_n, f_{n,s}, G_n)$ .
- **2** Estimate  $Q_0$  using the empirical distribution  $Q_n$ .
- **3** Set  $H_n := (G_n, Q_n)$ , and plug in estimated components:

 $V(f_n, H_n) - V(f_{n,s}, H_n) .$ 

#### Plug-in estimation

The bias of the plug-in estimator tends to zero at a rate slower than  $n^{-\frac{1}{2}}$ .



This is due to the fact that  $H_n$  is not targeted to the parameter of interest. Interestingly, the estimation of  $f_0$  and  $f_{0,s}$  does not contribute to the excess first-order bias (Williamson et al., 2023).

#### Debiased estimation

How can we recover  $n^{-\frac{1}{2}}$  asymptotics while using flexible nuisance estimators?

- 1 One-step estimator (Pfanzagl, 1982)
- 2 TMLE (van der Laan and Rubin, 2006)

Consider a parameter mapping  $P \mapsto \Psi(P)$  that is pathwise differentiable (that is, smooth) with gradient  $\phi$ . For any estimator  $\hat{P}_n$  of  $P_0$ , a first-order expansion, similar to a functional Taylor expansion, gives

 $\Psi(\hat{P}_n) - \Psi(P_0) = \mathbb{P}_n \phi(P_0) - \mathbb{P}_n \phi(\hat{P}_n) + R(\hat{P}_n, P_0) + (\mathbb{P}_n - P_0) \{\phi(\hat{P}_n) - \phi(P_0)\}$ 

- : Linear term, determines first-order behavior
- : Bias term
- Second-order remainder term
- : Empirical process term

Under some conditions, we can expect the one-step estimator  $\Psi(\hat{P}_n) + \mathbb{P}_n \phi(\hat{P}_n)$  to behave approximately like  $\mathbb{P}_n \phi(P_0)$ .

In general, there are multiple possible one-step estimators:

- 'direct' debias: using the gradient of  $P \mapsto V(f_0, H_P)$
- 'indirect' debias: construct targeted estimator  $H_n^*$  of  $H_0$  using the gradient of  $P \mapsto H_P$ , then construct estimator  $V(f_n, H_n^*)$

The gradient of  $P \mapsto H_P$ , evaluated at the point  $(x_0, t_0)$ , is given by

$$(x, y, \delta) \mapsto \mathbb{1}(x \le x_0) \left[ G_0(t_0 \mid x) + S_0(t_0 \mid x) \left\{ \frac{\delta \mathbb{1}_{[0,t]}(y)}{S_0(y \mid x) R_0(y \mid x)} - \int_0^{t \land y} \frac{L_0(du \mid x)}{S_0(u \mid x) R_0(u \mid x)} \right\} \right]$$

 $(G_0, S_0, L_0)$ : cdf, survival function, cumulative hazard of T given X.  $(R_0)$ : survival function of C given X.

$$\psi_{n,s}$$
: direct one-step  $\psi_{n,s}^*$ : indirect one-step

We have thus far avoided the question of how to estimate the oracle prediction functions  $f_0$  and  $f_{0,s}$ .

- ▶ Derivation of the form of  $(f_0, f_{0,s})$  must be handled on a case-by-case basis.
- ► For commonly used choices of W, doubly-robust estimation is possible consistent estimation of (f<sub>0</sub>, f<sub>0,s</sub>) is achieved by consistent estimation of either G<sub>0</sub> or R<sub>0</sub>.

## AUC, Brier score, MSE for $\tau$ -restricted survival time, . . .

- ► The oracle prediction function can be written as E<sub>P<sub>0</sub></sub>{h(T) | X = x} for a function h
- Use the doubly-robust pseudo-outcome regression approach of Rubin and van der Laan (2007)

#### C-index

- Oracle prediction function not available in closed form
- Estimation through direct optimization of  $f \mapsto V(f, H_n^*)$ .

#### Result:

- **1** When all nuisances are estimated well,  $\psi_{n,s}$  and  $\psi_{n,s}^*$  have identical first-order asymptotics.
  - $\blacktriangleright$  Letting  $\phi_0$  denote the efficient influence curve of the variable importance parameter, we have

$$\psi_{n,s} = \psi_{n,s}^* + o_P(n^{-\frac{1}{2}}) = \frac{1}{n} \sum_{i=1}^n \phi_0(X_i, Y_i, \Delta_i) + o_P(n^{-\frac{1}{2}}) .$$

► Therefore,  $n^{\frac{1}{2}}(\psi_{n,s}^* - \psi_{0,s}) \rightsquigarrow N(0, \sigma_{0,s}^2)$  with  $\sigma_{0,s}^2 := \operatorname{var}_{P_0} \{\phi_0(X, Y, \Delta)\}.$ 

**2** As long as either  $G_0$  or  $R_0$  is estimated consistently,  $\psi_{n,s}^*$  remains consistent, while  $\psi_{n,s}$  may fail to be.

$$\psi_{n,s}^* \xrightarrow{p} \psi_{0,s} \qquad \qquad \psi_{n,s} \xrightarrow{p} \psi_{0,s}$$

- Our proposed procedure enjoys doubly-robust consistency.
- Doubly-robust inference (confidence intervals and p-values) in this setting remains an open question (Benkeser et al., 2017).

## Cross-fitting

A standard regularity condition for asymptotic linearity is that  $(f_n, f_{n,s}, G_n, R_n)$  are not too complex. This is called a Donsker condition.

Cross-fitting can help us avoid this condition (Zheng and van der Laan, 2011; Chernozhukov et al., 2018).



#### Simulations: robustness

Scenario: conditional event distribution estimator misspecified

Empirical bias and variance near zero using

- indirect debiasing
- doubly-robust pseudo-outcome estimation of  $(f_0, f_{0,s})$



Oracle estimator: — Conditional surv. function -- Doubly-robust pseudo-outcome Debiasing: D Direct \* Indirect

#### Simulations: $n^{-\frac{1}{2}}$ -rate estimation and inference

**Scenario:** nuisances estimated using the global survival stacking ensemble learner (Wolock et al., 2024)

The cross-fitted estimator has

- second-order bias,
- > variance proportional to the nonparametric efficiency bound, and
- coverage near the nominal level.



Method: - Not cross-fit -- Cross-fit

## Variable importance in HVTN 702



#### Full cohort

- Sex assigned at birth is clearly an important predictor.
- Qualitative results are fairly stable across time horizons for AUC.

## Variable importance in HVTN 702



#### Variable importance and EHR data

Assessing variable importance can be more complicated when using EHR data.

#### Truncation-induced selection bias:

Patients enter the study after the initiating event and before the terminating event.

Example: Cancer patients who must undergo genetic testing after diagnosis to enter the study.



Truncation induces selection bias.

#### Limited outcome labeling:

Ascertainment of suicide death requires linkage with state mortality records; this information is missing for certain health systems.



Morenz, Wolock, Carone. "Debiased machine learning for counterfactual survival functionals based on left-truncated right-censored data." Wolock, Yan, Ning, Chen. "Transfer learning for model-free variable importance." (in progress)

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## Opportunities: inference under the null

Recall: We compare the predictiveness of

- *f*<sub>0</sub> ∈ *F*, an unrestricted class of prediction functions;
- *f*<sub>0,s</sub> ∈ *F*<sub>s</sub>, the class of prediction functions that does not use *X*<sub>s</sub>.

Suppose we wish to test the null hypothesis  $H_0: \psi_{0,s} = 0$ .







 $V(f_n, H_n^*)$  and  $V(f_{n,s}, H_n^*)$ have identical influence functions.

See Dai et al. (2022) and Hudson (2023) for work in this area.

CASCADIA study: How important are levels of neutralizing and binding antibodies, measured at baseline, for predicting risk of SARS-CoV-2 infection?



- Recruitment is by household
  - 1 Data units correlated
  - 2 Variable # of individuals per household
- ► Open questions:
  - 1 How to debias?
  - 2 How to make inference?

A nonparametric variable importance analysis can help provide insight into a given prediction task and inform future data collection practices.

Our proposed framework

- ▶ is nonparametric and algorithm-agnostic;
- accommodates censoring informed by measured covariates;
- encompasses commonly used predictiveness measures;
- provides doubly-robust estimation and calibrated statistical inference while allowing for flexible nuisance estimation.

Interesting and practically important work remains to be done, including

- improved procedures for inference under the null;
- efficient estimation and inference with correlated data.

See the  ${\tt survML}$  package on CRAN for implementation of the methods discussed today.

#### Other work

Estimating symptom duration following SARS-CoV-2 infection using current status data



Estimating causal effects from EHR data with underreported exposure



Wolock et al., "Investigating symptom duration using current status data: a case study of post-acute COVID-19 syndrome." Wolock et al., "Estimating causal effects from electronic health records data with underreported exposure." (in progress)

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#### Thank you for listening!

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